# An Empirical Assessment of the Efficiency of Trading Halts to Disseminate Price-Sensitive Information During the Opening Hours of a Stock Exchange. The Case of Euronext Brussels

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#### Abstract

This paper examines the effect of temporarily suspending the trading of exchange-listed individual stocks. It evaluates whether regulatory authorities can successfully use the mechanism of trading halts in forcing companies to disclose new and material information to the capital market. In contrast to previous studies which mainly concentrate on North-American stock markets, this study utilises a data set comprising of firms listed on Euronext Brussels. The results show that suspension is indeed an effective means of disseminating new information. Stock prices adjust completely and instantaneously to the new information released during trading suspensions. We also observe a significant increase in trading volume after the reinstatement of trading. On the other hand, we do not find any increase in stock return volatility around trading suspensions. Overall, the results confirm the efficacy of trading suspensions in disseminating new information.

JEL-Code: G10, G14, G18 Keywords: information dissemination, trading halt, trading suspension, market efficiency

#### **0.Introduction**

Do investors care about the quality of the financial markets in which they operate? Although an affirmative answer seems obvious and natural, only recently the so-called law and finance literature, which was initiated by the seminal papers of La Porta, Lopez-de-Silanes, Shleifer and Vishny (1997, 1998), investigated the relationship between a country's legal framework and its financial development. The law and finance literature offers strong empirical evidence on the importance of the legal environment (market integrity, investor protection) for the development of these markets and economic growth. La Porta e.a. (1997) show that a good legal environment expands the ability of companies to raise external finance through either debt or equity. In its recent proposal of the Market Abuse Directive, the European Commission as well stresses the importance of an adequate legal framework in order for companies to raise capital<sup>1</sup>: "High growth entreprises depend on the efficiency and transparency of financial markets in order to raise capital. Indeed, the smooth functioning of financial markets and public confidence in them are prerequisites for sustained economic

<sup>&</sup>lt;sup>1</sup> European Commission, Explanatory Memorandum, *Proposal for a Directive of the European Parliament and of the Council on Insider Dealing and Market Manipulation*, 30 May 2001, COM(2001) 281 final.

growth and wealth." The Forum of European Securities Commissions  $(FESCO)^2$  and the Federation of European Stock Exchanges (FESE) subscribe to that statement as well (see FESCO, 2000, 2001 and FESE, 2000, 2001).

Although FESE and the European Commission are both committed to highest levels of market quality, they disagree on the implementation. The European Commission supports the idea of a single central, administrative regulatory authority in each EU Member State, which should be independent of the exchanges. Because many exchanges demutualised fully, confess to a profit motive and have become listed companies themselves such as Stockholm, Deutsche Börse, Euronext and LSE, the European Commission argues that such exchanges cannot any longer be entrusted with regulation as they have a conflict of interest (Arlman, 2001a, 2001b). However, for a stock exchange to be successful in the long term, it must be able to show that it runs a fair and honest market for all participants, otherwise investors will avoid trading in such a market. Or as Arlman (2002) puts it: "Serving as a quality market is their most important brand." Instead of a central authority substituting the supervision of the exchanges, FESE (2001) stresses the complementary role of the exchanges and their experience and investment in detection and market surveillance.

As the European Commission defines market abuse as a situation in which investors "have used information which is not publicly available to their own advantage", the supervision of ongoing disclosure obligations, particularly to the ad-hoc disclosure of price-sensitive news, is of crucial importance in order to promote market integrity and market quality. In this respect, FESE (2001) particularly stresses the need for market knowledge and proximity and supports the allocation of regulatory powers 'closer to the market'. The alleged benefits of allocating the supervision to the exchange itself are: (a) familiarity with the trading system and its screening and filtering algorithms, (b) closer contact with market participants and (c) more timely actions when market irregularities are detected, all of which are crucial elements for installing trading halts for disseminating price-sensitive news. This empirical study adds to the above debate because it examines whether an exchange can add value by ensuring market integrity. In particular, it evaluates the use of trading halts on Euronext Brussels<sup>3</sup> to disseminate price-sensitive information during the opening hours of the stock exchange. Put

<sup>&</sup>lt;sup>2</sup> Now called the Committee of European Securities Regulators (CESR).

<sup>&</sup>lt;sup>3</sup> Formerly known as the Brussels Stock Exchange (BXS). In 2000 BXS, Amsterdam (Amsterdam Exchanges) and Paris (ParisBourseSBF) merged to the form the single securities market *Euronext*.

differently, it evaluates whether a supervision 'close to the market' can maintain an orderly market.

A trading suspension (also known as trading halt) represents a temporary interruption in official trading of an individual stock on a stock exchange<sup>4</sup>. Authorities usually adopt this regulatory measure to provide investors extra time to evaluate newly released information about a specific company. It is especially used when there may have been a breach of confidence in relation to inside information or market manipulation, in which case companies are required to disclose additional information. Therefore, trading suspensions are said to be a crucial regulatory and supervisory measure in order to maintain a fair and orderly market in which "all investors should have simultaneous access on a timely basis to the information they require to take their investment decisions (FESCO, 2001)."

The desirability of trading suspension is subject to debate among regulators, market participants and academics. Proponents of trading suspension argue that it provides traders extra time to evaluate newly released information so that no specific group of investors obtain an unduly advantage in stock trading. They also argue that stock prices become more informative, uncertainty is reduced and investors are protected from volatile price movements. On the other hand, critics argue that trading suspension simply delays stock price adjustments, imposes additional costs to investors who are deprived of trading opportunities and makes an exchange less attractive to investors. Ultimately, it is the supervisory authority that needs "to weigh the benefits of allowing continuous trading against the desirability of interposing processes which afford market users the opportunity to reassess a changed situation and to alter their orders accordingly<sup>5</sup>."

<sup>&</sup>lt;sup>4</sup> The term 'trading halt' is often used to refer to different kinds of regulatory measures. First, it has to be distinguished from a *circuit breaker* which involves a market-wide halt of trading of all stocks because of the movement of prices (or volumes) beyond pre-set parameters in order to reduce market volatility (e.g. rule 80B of the NYSE). Besides market-wide circuit breakers, restrictions on daily price variations of individual financial instruments also exist, such as the static and dynamic volatility interruptions on Euronext (see rule 4404/1 Euronext Rule Book and Euronext Instruction nr.4-01, Euronext Cash Market Trading Manual, Notice nr.2001-3807, 29 October 2001 and 2001-3840, 30 October 2001). Second, it has to be distinguished from *listing suspensions* when the supervisory authority decides to suspend the listing of a particular financial instrument until the situation of non-compliance with the continuing obligations arising from a listing, has been remedied. Finally, *delisting* refers to the permanent cancellation of the listing. The terms 'trading halt' and 'trading suspension' will be used interchangeably throughout the remainder of this paper and refers to suspensions related to the dissemination of price-sensitive information.

<sup>&</sup>lt;sup>5</sup> FESCO, *Standards for Regulated Markets* (99-FESCO-C), standard 7, paragraph 14.

This paper assesses the efficiency of trading halts to disseminate information among market participants on Euronext Brussels. Therefore, we examine the pattern of trading activity before and after the trading suspension in order to evaluate this regulatory policy measure. Moreover, based on detailed information provided by the stock exchange, the empirical analysis traces if the return behavior surrounding the trading halt is affected by the publicly announced reason for the suspension. This is of major importance, because section one shows that existing empirical studies mainly provide evidence for North-American stock markets, while European stock markets are barely investigated with regard to the efficiency of trading suspension. This current empirical research can therefore add new evidence with regard to the use of a regulatory measure on small stock exchanges by examining trading suspensions on Euronext Brussels. Moreover, examining Euronext Brussels offers an opportunity to evaluate the efficiency of trading halts on a order-driven market with an electronic automated trading system without any influence of market-makers or specialists on the effect of a trading suspension, as is the case, for instance, on the New York Stock Exchange (NYSE). While trading halts on the NYSE tend to protect specialists (Howe and Schlarbaum, 1986), trading suspensions on Euronext Brussels don't appear to protect any particular member or interest group, but are intended to protect investors in general.

The remainder of this paper is organized as follows. In the next section, we review the relevant literature. Section three explains the research design. It describes the data, the sample, the variables and the methodology. Empirical results are reported in section four. Conclusions and policy recommendations appear in the final section.

#### **1.Review of Literature**

Although several theoretical discussions on the use of a particular regulatory measure can be made (see supra), the answer to the question on the efficiency of securities regulation is mainly an empirical one. Several empirical studies investigated the suspension of trading of a particular stock, mainly on North-American stock markets. However, the empirical results on the efficiency of trading halts show mixed results.

Hopewell and Schwartz (1978) examine NYSE-initiated trading halts from 1974 till 1975. News suspensions lasted for 261 minutes or approximately four and one-half hours<sup>6</sup>. Delayed openings tended to last longer than intraday suspensions<sup>7</sup> (437 minutes compared to 149 minutes). Securities experience a relatively large price adjustment during a trading halt, and the longer the trading halt, the larger the price adjustment<sup>8</sup>. Their empirical results show that security prices adjust rapidly to the new information released during the trading halt. Post-trading halt abnormal return behavior shows that the price adjustment is almost complete by the end of the trading halt. However, a commission-paying investor could not earn abnormal trading profits. While prices react efficiently to significant new information, pre-trading halt anticipatory price behavior can be detected. Although their analysis cannot distinguish between the possible causes of this abnormal return behavior, Hopewell and Schwartz suggest insider trading as a possible explanation. A distinction between bull and bear markets does not result in different conclusions (Hopewell and Schwartz, 1976).

Apart from the stock exchanges, also the Securities and Exchange Commission (SEC) can halt trading in U.S. markets. These trading suspensions are often intended to force compliance with reporting and disclosure requirements to protect investors or to promote investor equity by ensuring that sufficient information is available to make rational, informed decisions. Compared to NYSE halts, SEC-initiated trading halts are less frequent (frequency of SEC-halt is one tenth of NYSE-halts). The average length of the SEC-trading halts is substantially longer: on average 12.2 weeks (Howe and Schlarbaum, 1986). It appears that SEC-trading halts have to be viewed rather as a disciplinary measure to force compliance with securities regulation. Howe and Schlarbaum find a substantial negative abnormal return over the trading halt period. After the trading halt stock prices keep declining in the weeks thereafter. Apparently, SEC-trading halts disclose unfavorable news, to which investors react very slowly. No anticipatory abnormal return behavior is found. The Ferris, Kumar and Wolfe (1992) study, which examines 40 SEC-initiated trading halt during 1959 and 1987, detects anticipatory price behavior as well as no complete price adjustment to new information released during the trading halt, especially for the bad news subsample.

<sup>&</sup>lt;sup>6</sup> News suspensions are initiated by a pending or actual news announcement which is deemed to have a significant impact on the market price of a security. The NYSE has another event triggering a trading halt: a substantial imbalance of buy and sell orders and are requested by specialists. These trading halts are called imbalance suspensions. See Hopewell and Schwartz (1978) for more details.

<sup>&</sup>lt;sup>7</sup> Trading halts that are initiated sometime after the opening trade.

<sup>&</sup>lt;sup>8</sup> Similar results are being reported by Schwartz (1976).

Other empirical studies focusing on major stock exchanges are Kryzanowski (1978, 1979) on Canadian stock exchanges and Kabir (1994) on the London Stock Exchange (LSE). Kryzanowski (1978) reports abnormal returns in the presuspension period as well as in the postsuspension period. He concludes that trading halts are not an effective mechanism to detect the exploitation of private information nor to disclose price-sensitive information during the suspension of trading. Examining a slightly different sample Kryzanowski (1979) distinguishes between good and bad new information. Both subsamples shows abnormal returns prior to the trading halts, but only the bad news subsample show abnormal negative returns in the postsuspension period. Only the disclosure of favourable news appears to be efficient. Examining the trading halts on the LSE, Kabir (1994) confirms the doubts on the efficiency of this mechanism in disseminating price-sensitive information. Anticipatory price behavior as well as abnormal returns in the month following the month of trading reinstalment are reported.

Where as the efficiency of trading halts on large stock exchanges as the NYSE, the Montreal Stock Exchange or the London Stock Exchange is doubtful, empirical results on smaller stock exchanges seem more promising. Examining the trading halts on the Stockholm Stock Exchange (Sweden) De Ridder (1990) concludes that this is an effective mechanism to disseminate new information. No abnormal return behavior is detected in the postsuspension period, so prices fully adjust over the trading halt period. Moreover, no anticipatory price behavior is found during the presuspension period indicating that insiders did not benefit systematically from their informational advantage. Analogous conclusions can be drawn regarding to the trading halts on the Amsterdam Stock Exchange (the Netherlands). The market authority of the Amsterdam Stock Exchange appears to be very efficient in utilizing this regulatory mechanism in order to disclose new information. Kabir (1992) detects no anticipation of any trading halt, nor any share price behavior in the postsuspension period indicating a possibility of abnormal profit-making. Share prices fully incorporate the information released during the trading halt. With regard to the Stock Exchange of Hong Kong, Wu (1998) shows that there are no abnormal profits in the postsuspension period. However, some anticipatory price behavior is detected.

While previous empirical studies focused on return behavior around trading halts recent empirical studies also examine volume and volatility patterns around the suspension of trading. Examining SEC-halts, Ferris, Kumer and Wolfe (1992) observe a higher stock return

volatility in the presuspension period as well as in the postsuspension period. Only several months later a significant decline of volatility is detected indicating that trading halts are not effective in immediately reducing return volatility. Analogous results are reported with regard to volume: higher than normal volume in the presuspension as well as in the postsuspension period. Only four weeks after the trading halt normal volume patterns reoccur. Also Kabir (1992) reports higher trading volume around trading halts on the Amsterdam Stock Exchange. In his study, trading volume is slightly higher than normal in the ten-day period before the trading halt, and significantly higher in the ten-day period after the suspension. Lee, Ready and Seguin (1994) report increased volume and volatility after the reinstalment of trading on the NYSE. Similar results are reported by Wu (1998) with regard to the Stock Exchange of Hong Kong. Kryzanowski and Nemiroff (1998) detect higher trading activity in the presuspension period on the Montreal Stock Exchange. Trading activity declines in the postsuspension period, but is still higher than in the period prior to the event window. Also increased volatility is reported in the presuspension and in the postsuspension period, but volatility only increases temporarily in the postsuspension period and decreases within five hours after the trading halt to its level of the pre-event window<sup>9</sup>.

This review of literature showed mixed results concerning the use of trading halts to disseminate information among market participants. The efficiency of this regulatory measure is doubtful on major stock markets as the NYSE, the Canadian market and the London Stock Exchange. The efficiency on smaller stock markets as Stockholm or Amsterdam is more promising. The remainder of this chapter analyzes whether the efficiency of trading suspensions on Euronext Brussels is high or low and whether these results confirm the empirical findings on the use of trading halts on small stock exchanges. Table 1 summarizes the review of literature.

<sup>&</sup>lt;sup>9</sup> Given the existence of specialists on the NYSE or the Montreal Stock Exchange, some empirical studies focus on the role of specialists around trading halts. Because the role of specialist is irrelevant to this empirical study on the Brussels Stock Exchange, it will not be explained in extenso. See for details King, Pownall and Waymire (1991) and Kryzanowski and Nemiroff (1998). These studies focus on the price discovery process during trading halts using specialist indications (sequential forecasts of the upper and lower bounds of the security's price at the resumption of trade). See King, Pownall and Waymire (1991, 518) for an example.

# Table 1. Overview of the empirical studies on trading halts

Study	Country	Final	Period examined in	Who	Anticipatory	Complete	Volume?	Volatility?
	_	sample of	study	initiated the	price	adjustment to		-
		news		trading	behaviour?	new information		
		suspensions		halt?		released during		
						trading halt?		
Schwartz (1976)	USA	242	Febr.'74 –Oct'74	NYSE <sup>a</sup>	-	-	-	-
Hopewell and Schwartz (1976)	USA	Bull: 201	Febr'74 – June'75	NYSE	yes	yes	-	-
		Bear: 300						
Hopewell and Schwartz (1978)	USA	501	Febr'74 – June'75	NYSE	yes	yes	-	-
Kryzanowski (1978)	Canada	34	Jan'67 –Dec'73	4 stock	yes	no	-	-
				exchanges				
Kryzanowski (1979)	Canada	Good: 43	Jan'67 –Dec'73	4 stock	yes	yes – good news	-	-
		Bad: 77		exchanges		no – bad news		
Howe and Schlarbaum (1986)	USA	49	Febr'59 – May'79	SEC	no	no	-	-
De Ridder (1990)	Sweden	137	Jan'80 – June'88	SSE	no	yes	-	-
Kabir (1992)	Netherlands	59	Jan'83 – March'89	ASE	no	yes	Increases	-
Ferris, Kumer and Wolfe (1992)	USA	40	Febr'59 – Oct'87	SEC	yes	no	Increases	Increases
Kabir (1994)	UK	83	Jan'70 – March'88	LSE	yes	no	-	-
Lee, Ready and Seguin (1994)	USA	518	<b>'88</b>	NYSE	-	no	Increases after	Increases after
							trading halt	trading halt
Wu (1998)	Hong-Kong	522	April'86-Dec.'93	SEHK	yes	yes	Increases	Increases
Kryzanowski and Nemiroff	Canada	412	March-Aug'88	MSE	-	yes	Increases	Increases
(1998)			May-Oct'89			-		temporarily
			Oct'90-March'91					

<sup>a</sup> NYSE: New York Stock Exchange; SEC: Securities and Exchange Commission; SSE: Sweden Stock Exchange; ASE: Amsterdam Stock Exchange; LSE: London Stock Exchange; MSE: Montreal Stock Exchange; SEHK: Stock Exchange of Hong Kong

# 2.Research Design 2.1.Data

The initial population consisted of all suspensions of common stocks of Belgian companies on Euronext Brussels from January 1992 through June 2000. This list totalled 210 trading halts. This means, on average, 2.06 trading halts per month and 0.10 trading halts per trading day or one trading halt every 10.1 trading days<sup>10</sup>. This list was provided by the Market Authority of Euronext Brussels. The list included for each trading halt: the company name, the date of the trading halt, the date of trading reinstatement, the last stock price before and the first stock price after the trading halt and finally, the detailed reason for suspending the trading<sup>11</sup>. When data was missing or was incomplete, additional data was collected from the leading Belgian financial newspaper De Financieel Economische Tijd. Share price data were collected from *Datastream*.

# 2.2.Sample description

The initial sample consisting of 210 trading halts was further reduced in several ways. First, 12 trading halts were deleted because these companies were delisted shortly after the suspension of trading, meaning no after-suspension market price of the stock was formed. These trading halts were mainly the result of a bankruptcy, a corporate reorganization ordered by court or a regulatory measure of the supervisory authorities for non-compliance of the disclosure regulation. Second, all related trading halts are left out of the sample. When the trading is suspended in one stock, trading in related companies is suspended as well. For example, when trading in Petrofina was suspended on 30 November 1998 because of the takeover bid by Total, trading in the stocks of the shareholders of Petrofina was suspended as well. In this way, seven other stocks were suspended: Tractebel, Electrabel, Sofina, NPM, Sidro, GBL and Electrafina. Therefore the trading halts of related companies are left out of the final sample. Thirdly, 31 observations were lost because of lack of data. These 31 trading halts include mainly very small, thinly traded companies for which no stock price data was available, such as Musson-Halanzy, Aiseau-Presle, Charbonnages du Hasard, Aurex or Roumanie-Société Générale des Sucreries. Finally, 22 observations were excluded because of

 <sup>&</sup>lt;sup>10</sup> Given 250 trading days in one year.
 <sup>11</sup> The author thanks Mr. V. Van Dessel and Mr. L. Delboo of Euronext Brussels for providing this data.

overlapping event periods as well as overlapping pre-event periods used for the estimation of parameters to calculate abnormal returns (see infra). For instance, *Ion Beam Applications (IBA)* was suspended on 15 February 1999 and again on 3 March 1999. Because the pre-event period of the second trading halt includes the first trading halt, the estimation of the parameters can be affected. Therefore, the second trading halt was left out of the sample.

In this way, the final sample consists out of 102 trading halts involving 72 companies. Of these companies, 48 (66.67%) were suspended only once during the sample period, while 24 companies were suspended more than once. Eighteen companies were suspended two times and six companies three times (see panel B in table 2). The average number of trading halts per company is 1.42 and the median is 1.00 (see panel A in table 2). Of the 102 trading halts from January 1992 through June 2000, 82 are single day suspensions (80.4%), while 20 are multiday suspensions (19.6%). The average suspension period is 2.34 days (see panel A of table 3). Panel B shows that 92% of all trading halts lasts two days or less.

Panel A. Number of trading halts								
Number of trading halts	102							
Number of companies	72							
Average number of trading halts per company	1.42							
Median number of trading halts per company	1.00							
Panel B. Distribution over number of trading halts per company								
Number of trading halts	Number of companies							
1	48							

 Table 2.Number of trading halts per company for the final sample

Table 3.	Single day	<sup>,</sup> versus multidav	v trading halts	of the	final	sample
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Panel A. Single versus multiday suspensions										
single day	82	80.4%								
multiday	20	19.6%								
average	2.34	min	1							
median	1	max	65							
Panel B. Dura	tion of suspen	sion (number of days)								
1	82	80.39%								
2	12	11.76%								
3	4	3.92%								
≥4	4	3.92%								

Legend: Final sample of 102 trading halts from January 1992 through June 2000

Table 4 give more descriptive statistics for the final sample of 102 trading halts. Panel A through C give the distribution of the trading halts per year, per month and per day of the week. Except for 1997, we notice that the number of trading halts increased substantially during the last five years. There is no specific pattern of trading halts throughout the year. Most trading halts occur in October, September and May, while the least trading halts occur in August and June. Although most of the trading halts occur on Thursday, no day-of-the-week pattern is present.

	Absolute number	Percentage	
Panel A.	Number of trading ha	lts per year	
1992	15	14.7%	
1993	9	8.8%	
1994	5	4.9%	
1995	5	4.9%	
1996	19	18.6%	
1997	8	7.8%	
1998	14	13.7%	
1999	18	17.6%	
2000	9	8.8%	
Panel B. N	Number of trading hal	ts per month	
Jan	8	7.8%	
Febr	7	6.9%	
March	9	8.8%	
April	8	7.8%	
May	10	9.8%	
June	3	2.9%	
July	7	6.9%	
Aug	6	5.9%	
Sept	11	10.8%	
Oct	16	15.7%	
Nov	9	8.8%	
Dec	8	7.8%	
Panel C. Num	ber of trading halts pe	r day of the week	
Monday	18	17.6%	
Tuesday	13	12.7%	
Wednesday	21	20.6%	
Thursday	30	29.4%	
Friday	20	19.6%	

Table 4. Descriptive statistics for the final sample

Legend: Final sample of 102 trading halts from January 1992 through June 2000

Because the list provided by the Market Authority of Euronext Brussels included detailed information on the reason for suspending the trading of each particular stock, each trading halt of the final sample was categorized according to a specific type of news. A detailed list of

these news categories is provided in appendix A. Table 5 gives a summary of this categorization. Most trading halts occurred because of the suspended company being a takeover target (34.3%), followed by restructuring (21.6%), corporate acquisitions (18.6%) and divestitures (13.7%).

Reason	Absolute number	%	
Corporate Acquisitions	19	18,6%	
Takeover targets	35	34,3%	
Financial Information	3	2,9%	
Divestitures	14	13,7%	
Restructuring except divestitures	21	21,6%	
Legal issues except bankruptcy	2	2,0%	
Bankruptcy	0	0,0%	
Miscellaneous	2	2,0%	
No news	6	4,9%	

*Table 5. Reason (news category) for the trading halt for the final sample* 

Legend: Final sample of 102 trading halts from January 1992 through June 2000

Price movements over the suspension period tend to be large indicating that significant pricesensitive information was released during the trading halt. The average raw return over the suspension period was 4.43% (median was 1.45%)<sup>12</sup>. The largest negative return is -47.31%, while the largest positive return is 53.68%. Similar figures are reported in other empirical studies, e.g. Schwartz (1976) reports positive returns of 52.9%, 58.8% and 75.0% and negative returns of -62.9% and -75.0%. The distribution of stock price returns over the suspension period is shown in figure 1. This figure shows that the majority of the trading halts cause a positive price movement over the suspension period, while 41 trading halts have a price movement equal to or smaller than zero. Furthermore, the figure shows that 33 trading halts experience a price movement of 2% (in absolute value). Half of the sample has a price movement within the 4%-range, while 28% of the suspensions have a price movement exceeding 10% (in absolute value).

12

<sup>&</sup>lt;sup>12</sup> Abnormal returns are reported in section three.



Figure 1.Frequency distribution of the average raw return over the suspension period for the final sample of 102 trading halts

#### 2.3.Description of the variables

In the rest of this paper the following terms are used:

 $R_{i, t}$  = the logarithmic return of stock i in period  $t^{13}$ 

 $R_{m,t}$  = the logarithmic market index return in period t

 $AR_{i,t}$  = abnormal return of stock i on day t of the estimation period

 $AR_{i,E}$  = abnormal return of stock i on the event day

 $R_{m,t}$  = market index return on day t of the estimation period

 $R_{m, E}$  = market index return on the event day

 $\overline{R}_m$  = average market index return during the estimation period

N = number of stocks in the sample

T = number of trading days in the estimation period

 $\hat{s}_i$  = estimated standard deviation of the abnormal return of stock i during the estimation period

 $SAR_{i, E}$  = standardized abnormal return of stock i on the event day

w = number of stocks in the sample with a positive abnormal return on the event date

#### 2.4. Methodology

To evaluate the efficiency of using trading halts to disseminate price-sensitive information among market participants, an event-time study is used to analyze the impact of the trading suspension. In this case the event is the suspension of the stock and therefore day 0 is defined as the day on which the trading halt occurs, while day -1 is the trading day immediately before the suspension day and day +1 is the day immediately after the suspension day. The return on day 0 is calculated between the last closing price before the trading halt to the first closing price after the trading halt, while the return of day +1 is calculated as the return between this first closing price and the next day's closing price<sup>14</sup>. While 80% of the trading halts are single day suspensions (see table 2), 20 trading halts out of 102 are multiday suspensions. In order to obtain comparable daily returns on event day 0, the multiday returns over the suspension period are scaled by the number of suspension days.

An event study examines if the average abnormal return on the event day is equal to zero (null hypothesis) versus an alternative hypothesis of a non-zero abnormal return:

$$\begin{cases} H_0: & AAR_E = 0\\ H_1: & AAR_E \neq 0 \end{cases}$$
[1]

The average abnormal return  $(AAR_E)$  on the event day is the aggregation of the individual stock abnormal returns aligned in event time:

$$AAR_E = \frac{1}{N} \sum_{i=1}^{N} AR_{i,E}$$
[2]

On the event day and on twenty trading days before and after the suspension, resulting in a 41-day event window<sup>15</sup>, abnormal returns are being calculated to examine returns behavior around the trading halt. Individual stock abnormal returns are measured as the difference between the realized or actual return on the event day ( $R_{i, t}$ ) and the expected return  $E[R_{i, t}]$ , which is the benchmark normal return in the absence of the event:

<sup>13</sup> The logarithmic return of a stock is calculated as:  $R_{i,t} = \ln\left(\frac{P_{i,t} + D_{i,t}}{P_{i,t-1}}\right)$ , where  $P_{i,t}$  is the *i*-stock closing price

on trading day t,  $D_{i, t}$  is the cash dividend paid on trading day t and  $P_{i, t-1}$  is the *i*-stock closing price on trading day t-1, all adjusted for all capital changes as stock splits and stock dividends.

<sup>&</sup>lt;sup>14</sup> Although more precise data on raw returns is available, i.c. the return between the last market price before the trading halt and the first market price after the trading halt, the corresponding data on market index returns lacked. Therefore, returns on day 0 are calculated from close to close.

<sup>&</sup>lt;sup>15</sup> The typical length of the event period ranges from 21 to 121 days for daily studies. See Peterson (1989).

$$AR_{i,t} = R_{i,t} - E[R_{i,t}]$$
[3]

Several methods exist to estimate the expected return of the stocks. In this study the market adjusted model and the market model are used. Moreover, the market model is adjusted to incorporate thin trading problems by using the Dimson (1979) methodology. These models are discussed in the next section.

#### 2.4.1.Calculating the benchmark expected return

The benchmark expected return for each individual stock depends on the model used:

 $E[R_{i,t}] = R_{m,t}$ , for the market-adjusted model,

 $E[R_{i,t}] = \hat{a}_i + \hat{b}_i \cdot R_{m,t}$ , for the market model,

 $E[R_{i,t}] = \hat{\alpha}_i^D + \hat{\beta}_i^D \cdot R_{m,t}$ , for the Dimson model, and

The expected return of a stock in the market-adjusted model is the current market index return. The market-adjusted abnormal return is thus equal to:

$$AR_{i,t} = R_{i,t} - R_{m,t} \tag{4}$$

This model uses no information from outside the event window to calculate abnormal returns during the event period.

Market model abnormal returns are calculated as:

$$AR_{i,t} = R_{i,t} - \left(\hat{a}_i + \hat{b}_i \cdot R_{m,t}\right)$$
[5]

where '^' denotes the OLS-estimates from the market model:

$$R_{i,t} = a_i + b_i \cdot R_{m,t} + e_{i,t}$$
[6]

with

 $R_{i,t}$  = the return of stock i in period t

 $R_{m,t}$  = the market index return in period t

 $a_i$ ,  $b_i$  = intercept and slope coefficient of the market model (stock-*i*-specific and time-independent parameters)

 $e_{i,t}$  = random disturbance term of the market model for stock i in period t

In order to calculate market model abnormal returns information from outside the event window is used. The parameters of the market model are estimated over a period from -21 to -140 trading days before the event day<sup>16</sup>.

When there is thin trading of stocks, the OLS-estimates of market model betas can be affected. Thin trading of stocks can reduce the measured correlation with the market index, and consequently the beta estimate: thinly traded stocks appear to have downward biased betas, while actively traded stocks have upward biased beta estimates. These biased beta estimates can cause biased abnormal returns and misspecified test statistics in event studies (Strong, 1992). Our sample consists of many thinly traded stocks. Therefore, the Dimson (1979)-method was used to adjust betas for the extent of thin trading<sup>17</sup>.

The estimation of the Dimson-beta consists of the aggregation of five estimated beta coefficients using two lead and two lag variables<sup>18</sup>:

$$\hat{\beta}_{i}^{D} = \sum_{k=-2}^{k=+2} \hat{b}_{k,i}$$
, or [7]

$$\hat{\beta}_{i}^{D} = \hat{b}_{-2,i} + \hat{b}_{-1,i} + \hat{b}_{0,i} + \hat{b}_{+1,i} + \hat{b}_{+2,i}$$
[8]

The variables  $\hat{b}_{k,i}$  with k = -2, -1, 0, +1, +2 are estimates of the slope coefficients in a multiple regression of the stock return in period t against the return on the market in periods t-2, t-1, 0, t+1 and t+2 (Dimson, 1979):

$$R_{i,t} = a_i + b_{-2,i} \cdot R_{m,t-2} + b_{-1,i} \cdot R_{m,t-1} + b_{0,i} \cdot R_{m,t} + b_{+1,i} \cdot R_{m,t+1} + b_{+2,i} \cdot R_{m,t+2} + w_{i,t}$$
[9]

While the OLS-estimation of beta uses the complete (-140,-21) estimation-window, the Dimson estimation uses an (-138,-23) estimation-window to allow for the two day leading and lagging. The abnormal returns are calculated as (Brown and Warner, 1985):

$$AR_{i,t} = R_{i,t} - \hat{\alpha}_i^D - \hat{\beta}_i^D \cdot R_{m,t}$$
[10]

with

$$\hat{\beta}_{i}^{D} = \sum_{k=-2}^{k=+2} \hat{b}_{k,i}$$
, and [11]

<sup>&</sup>lt;sup>16</sup> The typical length of the estimation period ranges from 100 to 300 days for daily studies. See Peterson (1989).

<sup>&</sup>lt;sup>17</sup> An alternative procedure is Scholes and Williams (1977). However, Fowler and Rorke (1983) show that the choice between Dimson and Scholes-Williams is equivalent.

<sup>&</sup>lt;sup>18</sup> Empirical studies use a large variety of leads and lags. For instance, Brown and Warner (1985) use k=-3,...,0,...,+3; Dimson and Marsh (1986) use k=-1,...,0,...,+5; Kabir (1994) uses k=-3,...,0,+1; O'Hanlon and Steel (1997) use k=-1,0,+1 and Ibbotson, Kaplan and Peterson (1997) use k=-1,0.

$$\hat{\alpha}_{i}^{D} = \frac{1}{116} \sum_{t=-138}^{t=-23} R_{i,t} - \hat{\beta}_{i}^{D} \frac{1}{116} \sum_{t=-138}^{t=-23} R_{m,t}$$
[12]

However, the use of procedures to correct for thin trading can be questioned. Brown and Warner (1985) show that there is no evidence that these procedures improve the specification or the power of the tests. Similar results were found by Dyckman, Philbrick and Stephan (1984). These findings were also reported by Reinganum (1982), Theobald (1983) and Cowan and Sergeant (1996)<sup>19</sup>. Strong (1992) points out that, although OLS beta estimates can bias the abnormal returns for an individual stock, these biases may average out to zero in the sample of the event study. Moreover, Bartholdy and Riding (1994) show that OLS even outperforms the use of alternative methods of beta estimation.

#### 2.4.2.Test statistics

The traditional test procedure assuming cross-sectional independence is the Patell (1976)-test<sup>20</sup>. This test statistic standardizes the abnormal return for each stock by its standard deviation. The resulting test statistic is given by equation  $[13]^{21}$ .

$$Z = \frac{\sum_{i=1}^{N} SAR_{i,E}}{\sqrt{\sum_{i=1}^{N} \frac{T_i - 2}{T_i - 4}}} \sim N(0,1)$$
[13]

with

$$SAR_{i,E} = \frac{AR_{i,E}}{\hat{s}_{i}\sqrt{1 + \frac{1}{T_{i}} + \frac{\left(R_{m,E} - \overline{R}_{m}\right)^{2}}{\sum_{t=-140}^{t=-21} \left(R_{m,t} - \overline{R}_{m}\right)^{2}}}}$$
[14]

However, traditional test statistics assume stable variances, meaning that there is no change in variance between the estimation period and the event period. Event-induced variance, on the

<sup>&</sup>lt;sup>19</sup> Strong (1992) reports similar results found by Gheyara and Boatsman (1980), Dodd and Warner (1983), Linn and McConnell (1983) and Dopuch et al. (1986).

<sup>&</sup>lt;sup>20</sup> Since the trading halts occur independent of each other and no event-date clustering is present, no correction for cross-sectional dependence is necessary. Moreover, Brown and Warner (1985) point out that dependence adjustment can be harmful compared to procedures which assume independence because tests assuming cross-sectional dependence are only half as powerful and usually not better specified than test assuming independence.

other hand, means that the variance during the event window exceeds the variance over the estimation period (Seiler, 2000). If the variance is underestimated, traditional test statistics will reject the null hypothesis too frequently, even when the average abnormal return is in fact zero (Brown and Warner, 1985; Boehmer, Musumeci and Poulsen, 1991). Several studies report indeed increases of the variance of returns when certain events occur. A parametric test that incorporates event-induced variance is offered by Boehmer, Musumeci and Poulsen (1991). This test improves the Patell-test by allowing the abnormal return variances to differ between the event and the estimation periods. They show that even a very small increase in variance is very problematic for the traditional tests. Their test statistic incorporates variance information from the estimation as well as the event window (Boehmer, Musumeci and Poulsen, 1991):

$$Z = \frac{\frac{1}{N} \sum_{i=1}^{N} SAR_{i,E}}{\sqrt{\frac{1}{N(N-1)} \sum_{i=1}^{N} \left(SAR_{i,E} - \sum_{i=1}^{N} \frac{SAR_{i,E}}{N}\right)^2}} \sim N(0,1)$$
[15]

with

$$SAR_{i,E} = \frac{AR_{i,E}}{\hat{s}_i \sqrt{1 + \frac{1}{T_i} + \frac{\left(R_{m,E} - \overline{R}_m\right)^2}{\sum_{t=-140}^{t=-21} \left(R_{m,t} - \overline{R}_m\right)^2}}}$$
[16]

The main disadvantage of parametric tests is that they are based on assumptions about the probability distribution of returns. Non-parametric tests do not depend on the assumption of normality. Because non-parametric tests do not use the return variance, these tests are more appropriate in case of event-induced variance. Two parametric tests are generally used: the sign test (see infra) and the rank test of Corrado (1989). Corrado (1989), Corrado and Zivney (1992) and Campbell and Wasley (1993) show that the rank test performs better than the traditional test statistics in case of event-induced variance. The rank test is given by equation [17] (Corrado, 1989 and Corrado and Zivney, 1992):

$$Z = \frac{1}{\sqrt{N}} \sum_{i=1}^{N} \frac{(U_{i,0} - 0.5)}{S_U} \sim N(0,1)$$
[17]

with

<sup>&</sup>lt;sup>21</sup> See section 1.3 for a description of the variables.

$$S_{U} = \sqrt{\frac{1}{161} \sum_{t=-140}^{+20} \left[ \frac{1}{\sqrt{N_{t}}} \sum_{i=1}^{N_{t}} \left( U_{i,t} - 0.5 \right) \right]^{2}}$$
[18]

$$U_{i,t} = \frac{K_{i,t}}{(M_i + 1)}$$
[19]

with  $K_{i,t}$  = rank (AR<sub>i,t</sub>), M<sub>i</sub> represents the number of non-missing abnormal returns for stock i and N<sub>t</sub> represents the number of nonmissing abnormal returns in the cross-section of N firms on day t in event time.

Besides event-induced variances, thin trading is another crucial problem for the event study test specification. Cowan and Sergeant (1996) point out that thinly traded stocks are characterized by numerous zero and large non-zero returns, causing non-normal return distributions. This causes traditional test statistics to be poorly specified (Campbell and Wasley, 1993). Similar results are reported by Maynes and Rumsey (1993) showing that the rank test is a good alternative for thinly traded stocks causing traditional tests to be misspecified. Cowan (1992) reports departures from normality (right skewness) causing parametric tests based on the normality assumption to be less well specified for Nasdaq stocks as compared to NYSE and AMEX stocks. Moreover, the rank test is also misspecified for Nasdaq stocks! However, the generalized sign test performs well for thinly traded stocks. The generalized sign test by Cowan (1992) is given in equation [22].

$$Z = \frac{(w - n\hat{p})}{\sqrt{n\hat{p}(1 - \hat{p})}} \sim N(0, 1)$$
[20]

where w represents the number of stocks in the sample with a positive abnormal return on the event date, p represents the fraction of positive abnormal returns expected under the null hypothesis, and

$$\hat{p} = \frac{1}{N} \sum_{i=1}^{N} \frac{1}{120} \sum_{t=-140}^{t=-21} \varphi_{i,t}$$
[21]

with  $\phi_{i,t} = 1$  when  $AR_{i,E} > 0$  and 0 otherwise<sup>22</sup>.

The poorly specification of the Patell-test is also confirmed by Cowan and Sergeant (1996). Their simulations show that the best test for thinly traded stocks with no increase of the return variance on the event date is either the rank or the generalized sign test. In case of an increase of the return variance on the event date, results are less clear. For lower-tailed tests the generalized sign test should be used. For upper-tailed tests the standardized cross-sectional test of Boehmer, Musumeci and Poulsen (1991) can be used, but it is not very powerful and it risks to be misspecified if the variance increase does not occur. An alternative is the generalized sign test, but it is misspecified in a few thinly traded samples. Results are summarized in table 6.

In the remainder of this paper we perform both parametric and non-parametric tests to determine statistical significance. The traditional Patell t-test assuming cross-sectional independence is performed first<sup>23</sup>. Next, we also use the generalised sign test of Cowan (1992) as a non-parametric test to test statistical significance of abnormal returns<sup>24</sup>.

*Table 6. Best replacement of the Patell-test in case of event-induced variance or thinly traded stocks* 

	Tickly traded stocks	Thinly traded stocks
No variance		
increase on	Patell-test	-Generalized sign test
event date		-Rank test
	1	
Variance	-Standardized cross-sectional test of	Lower-tailed tests:
increase on	Boehmer, et al. (1991)	-Generalized sign test
event date	-Rank test of Corrado (1989)	Upper-tailed tests:
	-Generalized sign test of Cowan (1992)	-Standardized cross-sectional test
		-Generalized sign test

Sources: Corrado (1989), Corrado and Zivney (1992), Cowan (1992), Campbell and Wasley (1993), Maynes and Rumsey (1993), Cowan and Sergeant (1996), Seiler (2000).

<sup>&</sup>lt;sup>22</sup> The difference between the generalized sign test and the traditional sign test is the value of p under the null hypothesis. While the traditional sign test uses a value of 0.50, the generalized sign test uses the fraction of positive returns in the estimation period, measured across N stocks and T days as value for p.

 $<sup>^{23}</sup>$  See equation [13].

<sup>&</sup>lt;sup>24</sup> See equation [20].

#### **3.**Empirical results

The empirical analysis of the efficiency of trading halts on Euronext Brussels starts with an examination of the abnormal returns in section 3.1. This analysis is completed by an analysis of the abnormal trading volume in section 3.2 and the volatility of the stock returns around the suspension in section 3.3.

# 3.1.Analysis of abnormal returns3.1.1.Complete sample

Because trading halts are a regulatory action designed to disseminate price-sensitive information among market participants, this section examines the process of the price adjustment to this new information before, during and after the suspension. Through an analysis of the valuation effects of the suspensions we can evaluate the effectiveness of trading halts as a regulatory measure. If Euronext Brussels were a semi-strong form informationally efficient stock market, the stock price would adjust instantaneously to the new information that was released during the trading halt. Moreover, semi-strong form efficiency would imply that there isn't any anticipatory price behavior in the presuspension period, nor any significant abnormal return behavior in the postsuspension period. In this section we examine (a) if trading halts are associated with an important release of information, (b) if there is any unusual return behavior before the suspension and (c) if there is a complete adjustment to new information released during the trading halt.

To examine the abnormal return behavior over the suspension period, a market adjusted model was used as the benchmark expected return (see equation [4]). Table 7 contains the results for the entire sample of 102 suspensions from 1992 through 2000. The mean abnormal return over the suspension period amounts 3,31%. The cumulative abnormal returns are visualized in figure 2.

In order to test whether the abnormal returns in the event window are significantly different from zero, a non-parametric test was used. For, as explained in section 2.4.2, the traditional t-test performs very poorly in case of thin trading or variance increase on event date. Because the sample in this study contains a large amount of thinly traded stocks, a generalized sign test

(see equations [20]) is used to test statistical significance of the abnormal returns. The traditional t-test is merely reported for the sake of completeness<sup>25</sup>.



Figure 2. Market adjusted mean CARs for the entire sample from 1992 to 2000

Note: CAR = market adjusted cumulative mean abnormal returns Sample size: n=102 suspensions

Table 7 shows that only the abnormal return over the trading suspension (day [0]) is significantly different from zero. It appears that there is no anticipatory price behavior in the presuspension period. The CAR over the presuspension period [-20, -1] is 0.45%, although insignificant. Although many previous studies show anticipatory price behavior prior to the trading halt (see e.g. Hopewell and Schwartz, 1978; Kryzanowski, 1979 or Kabir, 1994), our results are in line with the findings on other small stock exchanges as Stockholm (De Ridder, 1990) or Amsterdam (Kabir, 1992).

Once the trading suspension is over, share prices do not follow any particular pattern. Although the CAR in figure 2 shows a downward trend after the reinstalment of trading, this is not statistically significant. We can conclude that share prices instantaneously incorporate

<sup>&</sup>lt;sup>25</sup> However, the results of the t-test have to be interpreted with caution because under the conditions of this sample, the traditional t-test will reject the null hypothesis of a zero abnormal return too often while in reality there is no abnormal return present.

the new information released during the trading halt. Again, these results are in line with De Ridder (1990) and Kabir  $(1992)^{26}$ .

	Market adjusted (n = 102)										
	AR	CAR	t-test	p-value	Z-value Gen.Sign Test						
-20	-0,0017	-0,0017	-0,85	0,3977	-0,41						
-15	-0,0020	-0,0032	-1,35	0,1796	-0,21						
-10	0,0017	-0,0005	1,97	0,0516	-0,01						
-5	0,0025	0,0045	1,66	0,0995	1,38						
-4	-0,0019	0,0026	-1,19	0,2374	0,39						
-3	-0,0004	0,0022	-0,09	0,9296	-0,21						
-2	-0,0011	0,0011	0,81	0,4197	0,78						
-1	0,0034	0,0045	1,25	0,2154	0,39						
0	0,0331	0,0377	28,20**	0,0000	2,96*						
1	-0,0012	0,0364	0,22	0,8265	-0,01						
2	-0,0024	0,0341	-0,40	0,6918	0,19						
3	0,0038	0,0379	2,30*	0,0235	1,38						
4	-0,0010	0,0369	-0,64	0,5267	-1,40						
5	-0,0013	0,0356	-0,38	0,7017	1,18						
10	-0,0033	0,0276	-1,92	0,0572	-0,81						
15	0,0017	0,0253	0,54	0,5871	-0,01						
20	-0,0034	0,0132	-1,43	0,1555	-1,00						

Table 7. Market adjusted mean abnormal returns and cumulative abnormal returns for the entire sample from 1992 to 2000

Note:

*AR* = market adjusted mean abnormal return

CAR = market adjusted cumulative mean abnormal returns

Sample size: n=102 suspensions

t-test and p-value: test statistics for traditional t-test

*Z*-value: test statistic for generalized sign test

\*\* denotes significant at the 1% level

\* denotes significant at the 5% level

<sup>&</sup>lt;sup>26</sup> See table 1 for an overview of the abnormal return behavior on US markets.

#### 3.1.2. Robustness of the benchmark model

To examine the sensitivity of the above results to the choice of the benchmark model to calculate the abnormal returns, the analysis was repeated for the market model and the Dimson model. The parameters ( $\hat{\alpha}_i$  and  $\hat{\beta}_i$ ) of the market model are calculated over the 120 days estimation period, starting with day [-140] through day [-21]. Similar to Brown and Warner (1985) we excluded all observations that did not have at least 30 daily returns in the estimation period. In this way, the sample size was reduced to 82 observations. Because our sample contains many thinly traded stocks, the beta estimate of the market model will be biased downward (see section 2.4.1). Therefore, we use the Dimson-method to correct for thin trading, using two lead and two lag variables. Although the average beta obtained from the Dimson method (0.61) is much higher than the average beta from the OLS-estimation (0.14) it is still fairly low. The abnormal returns around the trading halts using the market model and the Dimson model as the benchmark expected return, are reported in appendix B. The results are similar to the results obtained from the market adjusted model. The mean abnormal return on day [0] is 3,43% and 3,45% for the market model and the Dimson model, respectively, compared to 3,31% using the market adjusted model. Again, the generalized sign test shows no significant abnormal returns prior to or after the trading halt.

Because the results are largely insensitive to the choice of the benchmark model, we use the market adjusted model as benchmark instead of the market model or the Dimson model in the rest of the chapter. First of all, the use of the latter would reduce the sample size from 102 to 82 observations. Secondly, the correct estimation of betas of thinly traded stocks is rather difficult. Thirdly, Fedenia, Hodder and Triantis (1994) show that the estimation of betas can seriously be distorted on stock exchanges, as Euronext Brussels, that are characterized by the presence of holding companies and equity cross-holdings.

#### 3.1.3. Subsamples based on news categories

Because the results, as shown in figure 2 and table 7, include the entire sample, it is difficult to interpret these price adjustments. Because the entire sample includes both positive and negative news, as well as different news categories (e.g. corporate acquisitions, restructuring or legal issues), aggregation across securities makes the results difficult to interpret because of

potential offsetting price impacts of the different subsamples. Therefore, we divide the total sample of trading halts in three subsamples according to the reason for the suspension. The first subsample contains 54 trading halts for news concerning corporate acquisitions and takeover targets<sup>27</sup>. This subsample is labelled "mergers and acquisitions". The second subsample includes all trading halts with regard to divestitures, as the sale of business segments, participations and spin-offs. This group includes 14 observations. Finally, a subsample of 21 trading halts is used for news related to other restructurings<sup>28</sup>. Although a finer partitioning of the data, according to the detailed scheme in appendix A, would be very useful, it is not possible because of the small sample size.





CAR = market adjusted cumulative mean abnormal returns Sample size: n=54 (Mergers and acquisitions), n=14 (Divestitures) and n=21 (Restructuring)

<sup>&</sup>lt;sup>27</sup> See appendix A for the different new categories.

<sup>&</sup>lt;sup>28</sup> News categories 41 to 45 in appendix A.

	Mergers and acquisitions (n=54)				Dive	stitures (n	n=14)			Restructuring (n=21)					
	AR	CAR	t-test	p-value	Z-value	AR	CAR	t-test	p-value	Z-value	AR	CAR	t-test	p-value	Z-value
-20	0,0016	0,0016	0,24	0,81	0,72	-0,0017	-0,0017	-0,53	0,60	-0,78	-0,0065	-0,0065	-0,92	0,37	-0,83
-15	-0,0036	-0,0035	-1,62	0,11	-0,37	0,0018	0,0125	0,62	0,54	0,29	-0,0031	-0,0074	-1,30	0,21	-0,83
-10	0,0022	-0,0045	1,91	0,06	0,45	-0,0028	0,0082	-0,21	0,84	-1,31	0,0068	0,0071	1,58	0,13	0,48
-5	0,0015	0,0047	0,80	0,43	0,99	0,0050	0,0103	1,76	0,10	0,83	-0,0013	0,0139	0,72	0,48	-0,39
-4	0,0000	0,0047	-0,74	0,46	0,99	0,0008	0,0110	0,91	0,38	0,83	-0,0046	0,0094	-1,00	0,33	-0,83
-3	0,0026	0,0074	0,99	0,33	1,27	0,0012	0,0122	0,14	0,89	-0,78	-0,0052	0,0042	-1,00	0,33	-1,70
-2	0,0047	0,0120	0,93	0,35	0,72	-0,0061	0,0062	0,35	0,73	0,29	0,0002	0,0044	1,12	0,28	0,92
-1	0,0027	0,0148	0,76	0,45	-0,37	0,0054	0,0115	1,89	0,08	-0,78	0,0017	0,0061	1,20	0,24	0,92
0	0,0803	0,0951	39,75**	0,00	3,72**	-0,0040	0,0075	3,24**	0,01	0,83	-0,0366	-0,0304	-1,69	0,11	0,48
1	-0,0048	0,0903	-1,84	0,07	0,17	0,0136	0,0211	2,59*	0,02	0,29	-0,0079	-0,0384	-0,30	0,77	-0,83
2	-0,0019	0,0884	-0,62	0,54	0,99	-0,0042	0,0169	-1,09	0,30	-1,31	-0,0039	-0,0423	-0,13	0,89	-0,39
3	0,0080	0,0964	3,41**	0,00	2,36*	0,0033	0,0202	-1,70	0,11	-0,24	-0,0145	-0,0568	-1,40	0,18	-1,26
4	0,0031	0,0995	0,34	0,73	0,45	-0,0029	0,0174	-0,71	0,49	-1,31	-0,0069	-0,0637	-1,93	0,07	-1,70
5	0,0011	0,1006	0,02	0,98	0,17	-0,0037	0,0137	-1,03	0,32	-0,24	0,0006	-0,0631	0,72	0,48	3,11*
10	-0,0056	0,0861	-2,00	0,05	-0,92	-0,0032	0,0195	-1,22	0,25	-1,31	0,0001	-0,0604	0,07	0,94	0,48
15	0,0014	0,0832	-0,13	0,90	-0,10	0,0002	0,0219	0,37	0,72	0,29	0,0077	-0,0403	0,87	0,39	0,48
20	-0,0039	0,0746	-1,59	0,12	-1,73	0,0010	0,0089	0,45	0,66	0,83	-0,0044	-0,0500	-1,12	0,28	-0,39

Table 8. Abnormal returns and cumulative abnormal returns over the event window for three subsamples

Legend: n = number of trading halts in the sample; AR = market adjusted mean abnormal return; CAR = market adjusted cumulative mean abnormal return; t-test & p-value = test statistic resp. p-value for the traditional t-test; Z-value = test-statistic for the generalized sign test; \*\* denotes significant at the 1% level and \* denotes significant at the 5% level

The results for the three subsamples are reported in table 8. Figure 3 contains a graphic representation of the CARs. None of the three subsamples shows any anticipatory price behavior. It appears that there is not any information leakage to the market with regard to mergers and acquisitions, divestitures or restructuring plans. Over the suspension period the mean abnormal return of the mergers and acquisitions subsample is 8,03%, which is significant at the 1% level. The mean abnormal return for the divestitures and restructuring subsamples are -0,4% and -3,66% respectively, although not statistically significant. Notice that figure 3 shows that the mergers and acquisitions subsample have, an average, a positive price impact, while the restructuring subsample has a negative price impact. If one compares figure 3 and figure 2, it is clear that the mergers and acquisitions subsample dominates the results of the total sample. Similar findings are reported by De Ridder (1990) for the Swedish stock market<sup>29</sup>. Furthermore, table 8 shows that there is no significant abnormal return behavior after the reinstalment of trading. It appears that stock prices adjust completely to the new information released during the trading suspension.

#### 3.1.4. Price impact over the span of the suspension period

Regardless of the sign of the price movement (positive or negative), figure 4 represents the magnitude of the abnormal returns over the event window [-20, +20]. It is clear that a trading halt is associated with the release of important price-sensitive information, resulting in an abnormal return over 8% (in absolute value) on day [0]. This is not surprisingly because the trading halts on Euronext Brussels are generally associated with the release of nonroutine and extremely price-sensitive information as mergers and acquisitions or restructurings. Routine announcements of earnings or dividends are in general not released during a trading halt. Only three cases out of the total sample of 102 trading halts concern the release of financial information as earnings or dividend announcements. Similar results are reported by King, Pownall and Waymire (1991) for the US: 79.3% of their sample is related to disclosures about corporate takeovers and leveraged buyouts, which cannot be predicted by investors, but which have large price impacts.

<sup>&</sup>lt;sup>29</sup> Compare figures 3 and 4 in De Ridder (1990). His subsample of mergers and acquisitions contains 70 trading halts out of a total sample of 137 observations.



Figure 4. Magnitude of the abnormal returns over the event window

ABS(AR) = absolute value of the market adjusted mean abnormal returns Sample size: n=89 (Mergers and acquisitions, divestitures and restructuring)

Figure 5. Median CARs for the three subsamples based on the reason for the suspension



Note:

median CAR = market adjusted cumulative median abnormal returns Sample size: n=54 (Mergers and acquisitions), n=14 (Divestitures) and n=21 (Restructuring)

#### **3.1.5.**The impact of outliers

To test the impact of outliers on the mean abnormal returns, median abnormal returns are calculated as well. The median abnormal return for the three subsamples are 3.11%, 0.19% and 0.16% for the mergers and acquisitions, divestitures and restructuring subsamples, respectively. In general, the conclusions of the median abnormal return analysis are similar to the mean abnormal returns. No anticipatory stock price behavior and complete price adjustment over the trading suspension (see table 9 and figure 5).

	Mergers and acquisitions			]	Divestitures	5	Restructuring		
	AR	p-value	CAR	AR	p-value	CAR	AR	p-value	CAR
-20	0,0003	0,4207	0,0003	-0,0042	0,4263	-0,0042	-0,0030	0,1349	-0,0030
-15	-0,0003	0,2025	-0,0046	0,0005	0,6698	-0,0033	-0,0040	0,1491	-0,0018
-10	0,0000	0,5156	-0,0062	-0,0039	0,2412	-0,0021	0,0025	0,3377	0,0061
-5	0,0016	0,3263	-0,0035	0,0023	0,4631	-0,0024	-0,0001	0,9861	0,0117
-4	0,0015	0,7962	-0,0020	0,0025	0,7064	0,0001	-0,0001	0,4340	0,0116
-3	0,0014	0,2056	-0,0006	-0,0003	0,9749	-0,0002	-0,0048	0,0885	0,0068
-2	0,0003	0,3705	-0,0002	0,0001	0,9515	-0,0001	0,0013	0,8649	0,0081
-1	-0,0002	0,6512	-0,0005	-0,0059	0,8077	-0,0060	0,0022	0,4041	0,0103
0	0,0311**	0,0000	0,0306	0,0019	0,7609	-0,0041	0,0016	0,5315	0,0119
1	-0,0001	0,4409	0,0305	0,0006	0,5830	-0,0035	-0,0019	0,5663	0,0100
2	0,0005	0,7664	0,0310	-0,0036	0,0580	-0,0070	-0,0002	0,7281	0,0099
3	0,0050**	0,0053	0,0360	-0,0007	0,7775	-0,0078	-0,0021	0,1790	0,0078
4	0,0001	0,7962	0,0361	-0,0010	0,3151	-0,0088	-0,0050	0,2722	0,0028
5	-0,0001	0,8228	0,0360	-0,0018	0,3792	-0,0105	0,0034*	0,0290	0,0062
10	-0,0009	0,1643	0,0352	-0,0035	0,4263	-0,0097	0,0013	0,6091	0,0111
15	-0,0001	0,6920	0,0323	-0,0003	0,8260	-0,0139	0,0016	0,3133	0,0133
20	-0,0028*	0,0453	0,0285	0,0013	0,5830	-0,0214	-0,0003	0,3754	0,0115

Table 9.Median abnormal returns and cumulative abnormal returns over the event window *for three subsamples* 

Note:

*AR* = market adjusted median abnormal return

*CAR* = market adjusted cumulative median abnormal returns

Sample size: n=54 (Mergers and acquisitions), n=14 (Divestitures) and n=21 (Restructuring) *p*-value: test statistics for the Wilcoxon signed rank test

denotes significant at the 1% level

denotes significant at the 5% level

#### **3.1.6.Good news versus bad news subsamples**

Besides subsamples based on the news categories of appendix A, another two subsamples are formed: a good news and a bad news subsample. In order to categorize a trading halt in one of the two subsamples, the tick sign test of Kraus and Stoll (1972), Hopewell and Schwartz (1976, 1978) and King, Pownall and Waymire (1991) was used. The tick sign is the sign of the price movement over the span of the suspension. This method permits the classification of those securities experiencing favorable and unfavorable developments such as good and bad news (Hopewell and Schwartz, 1976). If the return was positive, the trading halt was classified as good news (61 observations); if it was negative, it was classified as bad news (33 observations)<sup>30</sup>. The mean abnormal returns and CARs for the good and bad news subsamples are reported in table 10 and visualized in figure 6.





CAR = market adjusted cumulative mean abnormal returns

Sample size: n=33 (bad news) and n=61 (good news)

Again, both subsamples are in line with the predictions of a semi-strong form informationally efficient stock market. No anticipatory price behavior is detected and a complete and instantaneous price adjustment over the trading suspension is observed. The mean abnormal return for the good news subsample is 8.42%, while the mean abnormal return for the bad news subsample is -4.18%. The CARs for the good news subsample remain stable in the

postsuspension period, while the CARs for the bad news sample show a downward trend, which is, however, not statistically significant. In contrast to Kryzanowski (1979) and Howe and Schlarbaum (1986) we do not find lags and frictions in the downward adjustment of security prices to the release of unfavourable information during a trading suspension.

		Bad nev	vs sample	(n = 33)		Good news sample (n=61)				
	AR	CAR	t-test	p-value	Z-value	AR	CAR	t-test	p-value	Z-value
-20	0,0047	0,0047	1,17	0,250	1,66	-0,0046	-0,0046	-1,63	0,108	-1,62
-15	0,0018	0,0011	0,74	0,467	0,97	-0,0040	-0,0040	-2,38*	0,021	-0,85
-10	-0,0032	-0,0034	-0,06	0,955	-1,13	0,0036	-0,0026	2,19*	0,032	0,18
-5	-0,0012	0,0012	-0,15	0,883	0,27	0,0035	0,0038	2,14*	0,037	1,71
-4	-0,0030	-0,0018	-1,01	0,320	1,32	-0,0011	0,0027	-0,76	0,450	-0,34
-3	-0,0048	-0,0065	-1,35	0,187	-1,48	0,0021	0,0048	0,89	0,377	0,69
-2	0,0032	-0,0033	1,07	0,294	0,27	-0,0029	0,0019	0,07	0,942	0,18
-1	0,0002	-0,0031	-1,73	0,092	-0,43	0,0051	0,0070	2,48 <sup>*</sup>	0,016	0,43
0	-0,0418	-0,0449	-13,91**	0,000	-2,87**	0,0842	0,0912	45,88**	0,000	5,81**
1	-0,0123	-0,0572	-1,57	0,126	-0,08	0,0033	0,0945	0,80	0,428	0,69
2	-0,0054	-0,0626	-0,80	0,430	-0,78	0,0000	0,0945	0,13	0,893	1,71
3	0,0030	-0,0595	2,69*	0,011	1,66	0,0003	0,0947	-0,15	0,881	-0,08
4	-0,0090	-0,0686	-2,09*	0,045	-1,48	-0,0007	0,0940	0,05	0,960	-0,59
5	-0,0011	-0,0697	-0,11	0,916	1,32	-0,0028	0,0912	-0,29	0,773	0,43
10	-0,0008	-0,0784	0,34	0,733	-0,43	-0,0002	0,0834	-0,64	0,527	-0,34
15	0,0081	-0,0703	1,57	0,125	0,97	-0,0003	0,0808	-0,64	0,522	-0,59
20	-0,0030	-0,0855	-0,11	0,910	1,66	-0,0069	0,0772	-2,14*	0,036	-3,16**

Table 10. Abnormal and cumulative abnormal returns for the bad and good news sample

Note:

AR = market adjusted mean abnormal return

CAR = market adjusted cumulative mean abnormal returns

t-test and p-value: test statistics for traditional t-test

*Z*-value: test statistic for generalized sign test

\*\* denotes significant at the 1% level

\* denotes significant at the 5% level

#### 3.2. Analysis of abnormal trading volume patterns

Being a 'close to the market' supervisor, the Market Authority of Euronext Brussels monitors price as well as volume patterns of shares traded on the exchange. In some cases abnormal price or volume patterns indicate a possible unequal distribution of price-sensitive information among market participants and, in this way, a potential danger for insider trading. If abnormal volumes are detected and if there is a danger of unequal distribution of price-

<sup>&</sup>lt;sup>30</sup> Eight zero tick suspensions are excluded from the analysis.

sensitive information, then the Market Authority can halt trading in this share<sup>31</sup>. Besides analyzing the abnormal returns around trading halts on Euronext Brussels, the behaviour of abnormal trading volume around the suspensions is therefore investigated as well in this section. Moreover, as pointed by, for instance, Kabir (1992), Holthausen and Verrecchia (1990) and Stickel and Verrecchia (1994), a simultaneous price and volume study is necessary in order to assess the information content of an event more accurately. If trading halts show abnormal trading volumes, than these trading suspensions are likely to be associated with major information content.

Kabir (1992) reports higher than average trading volumes around suspensions, especially in the postsuspension period. The highest trading volume occurs on day [+1] and shows a decreasing trend from day [+1] through day [+10]. Also, Ferris, Kumar and Wolfe (1992) report higher abnormal trading volumes around the trading suspension. Their results show that trading volume returns to normal levels four weeks after the trading halt. An increase of trading volume after the trading suspension is found by Lee, Ready and Seguin (1994). Furthermore, the empirical results of Kryzanowski and Nemiroff (1998) and Wu (1998) show an increase of trading volume as well.

To analyze the abnormal trading volumes around the trading halt, we follow the methodology of Michaely, Thaler and Seguin (1994) and Wu (1998). First, the normal trading turnover for each stock was calculated over the estimation period from day [-100] to day [-21]. The normal trading turnover is defined as the number of traded shares by the number of outstanding shares of stock i:

$$TURN_{it} = \frac{VOLUME_{it}}{SHARES_i}$$
, with i = 1, 2, ..., N and t = -100, ..., -21 [22]

where  $VOLUME_{it}$  is the number of traded shares of stock i on date t, and  $SHARES_i$  is the number of outstanding shares of stock i. Next, on each trading day, the average trading turnover is calculated across firms:

$$TURN_t = \frac{1}{N} \sum_{i=1}^{N} TURN_{it}$$
, with t = -100, ..., -21 [23]

<sup>&</sup>lt;sup>31</sup> Interview with Mr. V. Van Dessel, President of the Market Authority of Euronext Brussels, in June 2000.

where N is the number of trading halts in the sample. Because of data availability the sample size was reduced from 102 trading halts to 61 trading halts. The average trading turnover is then calculated across all days in the estimation period [-100, -21]:

$$\overline{TURN} = \frac{1}{80} \sum_{t=-100}^{t=-21} TURN_t$$
[24]

Finally, the abnormal trading volume, measured as abnormal trading turnover, can be calculated over the event window  $[-20, +20]^{32}$ :

$$AV_E = \frac{TURN_E}{\overline{TURN}}, \text{ with } E = -20, \dots, +20$$
[25]

<b>Event period</b>	Abnormal trading turnover	t-statistic	p-value
-20	0,87	-0,40	0,6930
-15	1,21	0,68	0,5005
-10	0,89	-0,36	0,7220
-5	0,90	-0,33	0,7452
-4	1,08	0,25	0,8020
-3	1,19	0,60	0,5524
-2	1,32	1,00	0,3229
-1	1,16	0,51	0,6097
1	6,32	16,76**	0,0000
2	3,70	8,51**	0,0000
3	3,27	7,15**	0,0000
4	2,90	5,98**	0,0000
5	1,93	2,93**	0,0048
10	1,57	1,81	0,0757
15	1,40	1,25	0,2150
20	1,22	0,69	0,4902

Table 11. Abnormal trading volume patterns around trading halts on Euronext Brussels

Note:

Abnormal trading volume is measured as the abnormal trading turnover, i.e. the ratio between the daily trading turnover in the event period [-20, +20] and the daily average trading turnover across firms and across trading days in the estimation period [-100, -21]; Sample size: n=61 trading halts in the period 1992-2000 t-test and p-value: test statistics for t-test

\*\*\* denotes significant at the 1% level

\* denotes significant at the 5% level

<sup>32</sup> Note that the standard deviation can be calculated as:

$$STDEV_t = \frac{1}{79} \sum_{t=-100}^{t=-21} (AV_t - \overline{AV})^2$$
, with  $\overline{AV} = \frac{1}{80} \sum_{t=-100}^{t=-21} AV_t$ .

The abnormal trading volume is reported in table 11 and figure 7. Before the trading suspension no abnormal trading volume pattern is present. On the first trading day after the suspension the average daily trading turnover is six times as high as normal (significant at a 1% level). On day [+2] and [+3] the abnormal daily trading turnover is 3.70 and 3.27 (t-values are 8.51 and 7.15 respectively). Table 11 shows that abnormal volumes are found during the first five trading days after the suspension. Figure 9 clearly shows that the trading volume has a decreasing trend from day [+1] to day [+20]. It appears that the trading volume returns to its normal levels after ten trading days. A similar volume pattern is reported by Wu (1998). The increase of the trading volume during the first trading days after the suspension confirms the findings from the abnormal return behavior, indicating that trading halts are associated with an important release of information. Moreover, our results confirm the results of prior empirical studies as Kabir (1992), Ferris, Kumer and Wolfe 1992), Lee, Ready and Seguin (1994), Wu (1998) and Kryzanowski and Nemiroff (1998).

Figure 7. Abnormal trading volume pattern around trading halts



3.3.Analysis of stock return volatility

Besides analyzing abnormal return and trading volume behavior around trading halts, recent empirical studies also examine stock return volatility around suspensions (see e.g. Lee, Ready and Seguin, 1994 and Wu, 1998). For, this is a parameter which can be of interest for supervisory bodies in order to install a trading halt or not. This parameter is closely related to

the objectives of circuit breakers. This section investigates the impact of trading halts on stock return volatility. In fact, it is analyzed if a sudden information flux causes abnormal volatility around the trading halt. The stock price volatility is measured as the variance of daily stock returns. To obtain a benchmark estimate of normal volatility, the variance of daily returns over the historical period [-140, -81] was calculated<sup>33</sup>. Analogously, the variances for the complete suspension period [-20, +20], for the presuspension period [-20, -1] and for the postsuspension period [+1, +20] were calculated.

Skinner (1989) shows that the median is more representative of the true change in volatility than the mean. Therefore, table 12 tests whether the median variance around the suspension increases compared to the median historical variance (VAR hist). It appears that the median variance of the complete event period [-20, +20] is about twice that of the historical variance. To test if these medians are significantly different from each other, a Wilcoxon signed rank test was used. The Z-score for the Wilcoxon signed rank test that the median variance is the same in the two periods is -4.30, which is significant at the 1% level. This means that the variance in the event period [-20, +20] is higher than the historical variance. However, the higher variance in the event period is solely due to the large price jump over the trading halt. This can be seen when the event window is broken up in a presuspension and a postsuspension period. Although the variance of the presuspension period (VAR pre) slightly increases compared to the historical variance (VAR hist) and slightly declines in the postsuspension period (VAR post compared to VAR pre), the Z-values of the Wilcoxon signed rank test are insignificant. This means that the variances are not significantly different in the different periods (see table 12). Similar results are found if we use abnormal instead of raw returns.

Therefore, one can conclude that the volatility does not increase prior to or after the instalment of a trading halt. This evidence contradicts the results of Ferris, Kumar and Wolfe (1992) and Lee, Ready and Seguin (1994) for US markets and Wu (1998) for the stock exchange of Hong Kong. This is an important finding because it implies that the Market Authority does not have to worry about increasing volatility around suspensions. Therefore,

<sup>&</sup>lt;sup>33</sup> Similar to Ferris, Kumar and Wolfe (1992) a conservative choice of the historical period to calculate the normal volatility was used, by opting for a distant period [-140, -81] compared to the event period [-20, +20]. Ferris, Kumar and Wolfe (1992) use a historical period [-210, -151] compared to an event period [-60, +60].

the Market Authority should not focus on volatility as a parameter for their regulatory policy towards the dissemination of price-sensitive information. Our findings are especially interesting because the key issue in the decision by the Market Authority to halt trading is the expected price impact of the news, not the expected volatility<sup>34</sup>.

	RA	W RETUR	NS	ABNORMAL RETURNS			
	MEDIAN	Z-value	p-value	MEDIAN	Z-value	p-value	
VAR hist (-140, -81)	0,000269	-4,30**	0,0000	0,000237	-4,42**	0,0000	
VAR susp (-20,+20)	0,000598			0,000485			
VAR hist (-140, -81)	0,000269	-0,30	0,3837	0,000237	-0,12	0,4522	
VAR pre (-20,-1)	0,000280			0,000238			
VAR pre (-20,-1)	0,000280	-0,19	0,4248	0,000238	-0,70	0,2411	
VAR post (+1, +20)	0,000277			0,000256			
VAR hist (-140, -81)	0,000269	-0,43	0,3336	0,000237	-1,22	0,1111	
VAR post (+1, +20)	0,000277			0,000256			

Table 12. Volatility patterns around trading suspensions

Note:

VAR hist refers to the stock return variance calculated over the estimation period [-140, -81]; VAR susp refers to the stock return variance calculated over the event period [-20, +20]; VAR pre refers to the stock return variance calculated over the presuspension period [-20, -1] and VAR post refers to the stock return variance calculated over the postsuspension period [+1, +20]

Sample size: n=82 trading halts in the period 1992-2000

Z-value and p-value: test statistics for the Wilcoxon signed rank test that the variance is the same in the two periods

Abnormal returns are calculated using a market adjusted model

\*\*\* denotes significant at the 1% level

\* denotes significant at the 5% level

#### 4.Conclusions

Because investors care about the quality of the financial markets, it is of the upmost importance for the growth and development of European financial markets as a financing source for companies to exhibit the highest levels of market quality and market integrity. This is especially the case with respect to the ad-hoc disclosure of price-sensitive information during the opening hours of the stock exchange. This paper empirically examines trading

<sup>&</sup>lt;sup>34</sup> Interview with Mr. V. Van Dessel, President of the Market Authority of Euronext Brussels, in June 2000.

suspensions on the Euronext Brussels. The study is of particular interest because of three reasons.

First, these suspensions occur in order to compel firms to disclose new information to the market. This is different from suspensions of stock trading at times of imbalances in buy or sell orders and extreme volatility. In particular, this paper analyzed the use of trading halts to disseminate new information. If there arises a situation of unequal distribution of price-sensitive information among market participants because of an information leakage prior to a press release and there is a danger for insider trading, the Market Authority of Euronext Brussels can halt trading in the shares of the company at issue. This paper examined the efficiency of trading halts to disseminate price-sensitive information among market participants because information among market participants.

Secondly, prior studies show conflicting results with regard to the effectiveness of trading suspensions. Both efficient and inefficient stock price adjustments are documented. Although previous empirical studies on the NYSE, the Canadian stock market and the London Stock Exchange show mixed results on the efficiency of trading halts, the use of trading suspensions on smaller stock exchanges as Stockholm or Amsterdam seemed more promising. This paper analyzes whether the empirical results of Euronext Brussels are in line with these smaller stock exchanges.

Thirdly, this empirical study adds to the debate whether an exchange can add value by ensuring market integrity, compared to a single central, administrative regulatory authority which is promoted by the European Commission.

In this study, we focus on three different parameters: stock return, trading volume and return volatility to determine in a robust way the efficiency of trading suspensions. Examining the entire sample of trading halts we find that there is not any anticipatory or unusual return behavior before the suspension. Nor is there any significant abnormal return pattern in the postsuspension period, meaning that there is a complete and instantaneous adjustment to the new information released during the trading halt. It seems that the Market Authority was very effective concerning the correct timing to install a trading suspension. Moreover, the results also indicate the semi-strong form informationally efficiency of Euronext Brussels. Our results, furthermore, show that trading halts are associated with an important release of

information. This is not surprisingly because the trading halts on Euronext Brussels are generally associated with the release of nonroutine information, which is extremely price-sensitive, such as mergers and acquisitions or restructurings. On average, the magnitude of the abnormal return over the span of the trading halts was over 8%.

An analysis of three subsamples based on the news categories, shows similar results. On average, mergers and acquisitions have a positive price impact (+8.03%), while restructurings have a negative price impact (-3.66%). Our investigation shows that there is no anticipatory or unusual return behaviour before the suspension. Nor is there any significant abnormal return pattern in the post-suspension period. It indicates that there is a complete and instantaneous adjustment to the new information released during trading halts. It seems that the Exchange was successful concerning the correct timing to install a trading suspension as well as to reinstate trading. Also, an analysis of the good news and the bad news subsamples confirms these results.

Next, the abnormal trading volume was examined as well. We find an increase of the trading volume during the first five trading days of the trading halts. The trading volume pattern shows a decreasing trend returning to normal levels approximately ten trading days after the suspension. The abnormal volume analysis confirmed the results of the abnormal return analysis. Finally, the analysis of stock return volatility shows that volatility does not increase prior to or after the instalment of a trading halt.

Overall, our results confirm the efficiency of trading halts to disseminate price-sensitive information among market participants on Euronext Brussels. It appears that, in line with other small stock exchanges as Stockholm or Amsterdam, this regulatory action of the Market Authority on Euronext Brussels to disseminate price-sensitive new information among market participants is very efficient. One explanation for small stock exchanges to be more efficient in using trading halts as a regulatory measure is the 'closeness' of supervisory bodies of the exchange. For, the exchange is more familiar with its trading system and its screening devices, it has closer contact with market participants and it can react more timely when market irregularities are detected. With respect to the dissemination of price-sensitive information and the instalment of trading halts, the empirical results thus show that an allocation of regulatory powers 'close to the market' operates efficiently.

# Appendix A. News categories

# **1.Corporate Acquisitions**

10.Preliminary negotiations

11.Acquisition: formal proposal or agreement reached

12.Merger: formal proposal or agreement reached

# **2.Takeover targets** (suspended firm = target)

20.Rumor or preliminary disclosure of possible takeover

21.Formal takeover-bid

22.Substantial change in ownership, followed by formal bid or price maintenance

23.Proposal rejected or withdrawn or negotiations terminated

24.Leveraged buyouts

# **3.**Financial Information

30. Earnings announcements (negative news)

31.Management or analyst earnings forecast

32.Dividend announcement (dividend reductions)

# 4.Restructuring

40.Divestitures (sale of business segments, participations and spin-offs)

41. Announced intent to repurchase stock

42.Capital Structure Changes (stock/debt issues)

43.External restructuring plans (initiated plans or announced progress of restructuring)

44.Internal restructuring (e.g. new management, personnel cut)

45.Announcement of financial difficulties, corporate reorganization ordered by court (creditors) or liquidation of the company

# 5.Legal issues

50. Announced start of a legal action

51.Legal decisions in favor of the firm

- 52.Legal decisions against the firm
- 53.Legal decision concerning the firm

54. Announcement and status of bankruptcy

55.Involvement in criminal procedure

56.Other regulatory measure (e.g. by Banking and Finance Commission)

# 6.Miscellaneous

60.Important contract 61.Trading halt on other stock exchange

# 7.No news

70.Unknown or no news

71. Abnormal price behavior (insider trading, stock price manipulation)

# Appendix B. Results for the market model and Dimson model abnormal returns



Figure B.1. Market model mean CARs for the entire sample from 1992 to 2000

CAR = market model cumulative mean abnormal returns Sample size: n=82 suspensions

Figure B.2. Dimson model mean CARs for the entire sample from 1992 to 2000



CAR = Dimson model cumulative mean abnormal returns Sample size: n=82 suspensions

	Market model (n = 82)					Dimson model (n = 82)				
	AR	CAR	t-test	p-value	Z-value	AR	CAR	t-test	p-value	Z-value
-20	0,0007	0,0007	0,26	0,793	-0,31	-0,0001	-0,0001	-0,14	0,889	-0,54
-15	-0,0023	0,0011	-1,33	0,187	-1,64	-0,0010	0,0010	-1,03	0,305	-1,21
-10	0,0027	0,0071	1,92	0,058	1,02	0,0019	0,0056	1,91	0,059	-0,10
-5	0,0042	0,0195	2,42*	0,017	1,90	0,0042	0,0148	2,30*	0,024	2,55*
-4	-0,0009	0,0186	-0,74	0,460	-0,75	-0,0008	0,0140	-0,69	0,489	0,12
-3	0,0009	0,0195	0,64	0,521	-0,31	-0,0011	0,0130	-0,18	0,855	-0,99
-2	-0,0013	0,0182	1,07	0,288	0,35	-0,0010	0,0120	0,84	0,405	-0,54
-1	0,0047	0,0229	1,84	0,069	1,24	0,0034	0,0154	1,27	0,207	1,00
0	0,0343	0,0572	26,95**	0,000	2,34*	0,0345	0,0500	26,99**	0,000	2,77**
1	-0,0006	0,0566	1,11	0,271	0,80	-0,0017	0,0483	0,47	0,641	0,78
2	-0,0044	0,0522	-1,51	0,134	-0,97	-0,0040	0,0443	-1,29	0,200	-0,32
3	0,0006	0,0528	0,82	0,411	1,46	0,0009	0,0452	0,78	0,436	1,66
4	-0,0003	0,0526	0,30	0,766	0,13	-0,0020	0,0433	-0,62	0,538	-1,21
5	0,0009	0,0535	0,64	0,523	0,13	0,0006	0,0438	0,39	0,698	1,00
10	-0,0028	0,0436	-0,77	0,445	-0,53	-0,0029	0,0342	-1,25	0,213	-0,54
15	0,0009	0,0434	0,24	0,814	-0,75	0,0013	0,0341	0,11	0,910	-0,10
20	-0,0030	0,0311	-0,81	0,421	-1,86	-0,0037	0,0201	-1,06	0,292	-1,43

Table B.1. Market model and Dimson model mean ARs and CARs for the entire sample from 1992 to 2000.

Note:

AR = market model or Dimson model mean abnormal return

*CAR* = market model or Dimson model cumulative mean abnormal returns

*Sample size: n*=82 *suspensions* 

t-test and p-value: test statistics for traditional t-test

*Z*-value: test statistic for generalized sign test \*\* denotes significant at the 1% level

denotes significant at the 5% level

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